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ZBI ZENTRUM FÜR
BIOINFORMATIK

Score Matrices

Algorithms for Sequence Analysis

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(with material from Tobias Müller, Würzburg)

Summer 2021

Overview

Previously: Pairwise Sequence Alignment

- score maximization with general scoring schemes,
- Four variants: global, semiglobal, overlapping, local

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Previously: Pairwise Sequence Alignment

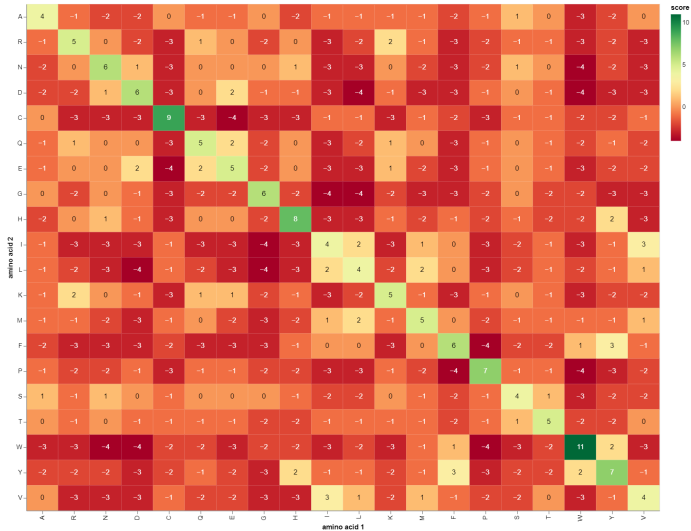
- score maximization with general scoring schemes,
- Four variants: global, semiglobal, overlapping, local

Today's Lecture: Score Matrices

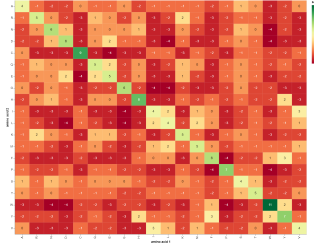
- Where do (families of) score matrices (like BLOSUM62) come from?
- Evolutionary distances (units PAM, PEM)
- Excursion: Time-continuous Markov processes, matrix exponentials
- Estimation of score matrices from alignments of different divergence times
- General vs. special purpose score matrices

Score Matrices for Comparing Proteins

Example: BLOSUM62 Scoring Matrix for Amino Acids

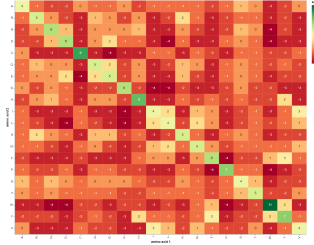


How Score Matrices are Obtained



- **Idea:** Physically and chemically similar pairs of amino acids have positive score, dissimilar pairs have negative score. Zero is a neutral value.
- However, who is to quantify “similarity”? Experts ?

How Score Matrices are Obtained



- **Idea:** Physically and chemically similar pairs of amino acids have positive score, dissimilar pairs have negative score. Zero is a neutral value.
- However, who is to quantify “similarity”? Experts ?
- Instead, use data-driven approach. (Today, you call this machine learning.)
- Observe the (relative) frequencies of amino acids in proteins.
- Observe the joint frequencies of amino acids in confirmed alignments:
Similar amino acids more often replace each other than dissimilar amino acids.

Score Matrices from Observed Alignments

Counting amino acid replacements in confirmed alignments

24.7% identity in 97 aa overlap; score: 109

```
ref  IFLHDNAPSHTARAVRDTLETLNWEVLPHAAAYSPDLAPSDYHLFASMGHALAEQRFDSYESVKKWLDEWFAAKDDEFYWRGIHKLPERWEKCVASDG
      .: .: .: .: .: .: .: .: .: .: .: .: .: .: .: .: .: .: .: .: .: .: .: .: .: .: .: .: .: .: .: .: .: .: .: .: .: .: .: .:
query VFQQDNDPKHTSLHVRSWFDRRFVDLLDWPSQSPDLNPIE-HLWEELERRLGGIRASNADAKFNQLPNAWKAIPMSVIHKLIDSMPRRCQAVIDANG
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Examples of pair counts

$$\#(D, E) = 4$$

$$\#(N, F) = 1$$

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Examples of pair counts

$$\#(D, E) = 4$$

$$\#(N, F) = 1$$

Assumption (for now)

All alignments used for counting have the same **degree of divergence** (evolutionary distance). Otherwise, they are not be comparable.

Markov Model of Protein Evolution

...	A	S	A	R	D	S	D	...
	↓	↓	↓	↓	↓	↓	↓	
...	D	S	D	A	A	S	D	...
	↓	↓	↓	↓	↓	↓	↓	
...	D	S	D	R	A	S	D	...
	↓	↓	↓	↓	↓	↓	↓	
...	A	E	D	A	D	S	D	...

Assumptions

- Replacement probabilities at any site depend only on the amino acid at that site and on transition probabilities, but not on the history (past) at that site.
- Sequences have average (typical) amino acid composition.
- Time-reversible process (direction of time arrow is irrelevant)

Markov Model of Protein Evolution

Model Parameters

- Amino acid frequencies $\pi = (\pi_i)_{i=1}^{20}$ (row vector)
- Conditional transition probabilities for a fixed time unit (“1 step”):

$$P = (P_{ij}) \quad (i = 1, \dots, 20, j = 1, \dots, 20)$$

with $\sum_{j=1}^{20} P_{ij} = 1$ for all i (rows sum to 1).

- One step of transitions must not change overall frequencies, i.e., π must be the/a **stationary distribution** for P :

$$\sum_{i=1}^{20} \pi_i \cdot P_{ij} = \pi_j \quad \text{or} \quad \pi \cdot P = \pi.$$

- Symmetric **joint** (or pair) frequencies $J_{ij} = \pi_i \cdot P_{ij} = \pi_j \cdot P_{ji} = J_{ji}$
(symmetry of J : **time-reversibility**)

Parameter Estimation

Procedure

- Estimate $J = (J_{ij})$ symmetrically by counting pairs of amino acids in alignments. Normalize, such that $\sum_{i,j} J_{ij} = 1$ (probability distribution over pairs), i.e., divide by total number N of observed pairs.
- Obtain π as marginals of J , i.e., $\pi_i = \sum_j J_{ij}$
- Obtain P by normalizing row sums of J to 1.

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Problem

- Procedure above assumes that all observed alignments have the same divergence time (“one step”).
- We will generalize this in a moment; for now, stick with the assumption.

Derivation of Score Matrix

Log-Odds Scores

- Compare the **observed** joint frequencies in real alignments with the **expected** pair frequencies based on amino acid frequencies:

$$\text{Quotient or enrichment or odds: } J_{ij}/(\pi_i \cdot \pi_j)$$

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- Scale and round to integer:

$$\text{Score } S_{ij} = \text{rd} \left(2 \log_2 \left(\frac{J_{ij}}{\pi_i \cdot \pi_j} \right) \right) \quad [\text{half-bits}]$$

Score Matrices of Different Divergence Times

Remember

Amino acid score matrices are rounded scaled log-odds scores, comparing observed joint frequencies with expected pair frequencies from marginals.

$$\text{Score } S_{ij} = \text{rd}\left(3 \log_2 \left(\frac{J_{ij}}{\pi_i \cdot \pi_j}\right)\right) \quad [\text{third-bits}]$$

Score Matrices of Different Divergence Times

Remember

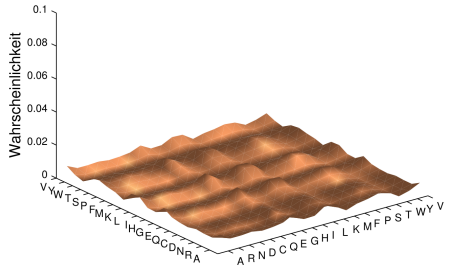
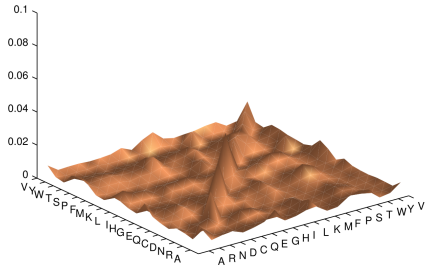
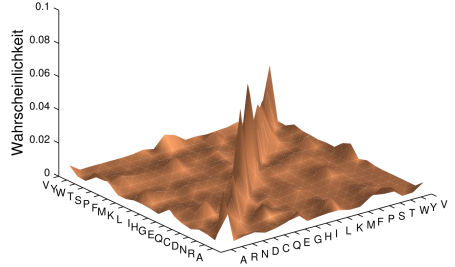
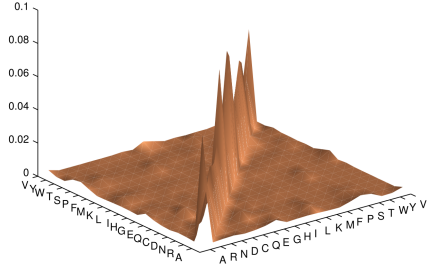
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Problem of Different Divergence Times: J_{ij} are mixed

VCKITPHSSNKSPYDGVYGTSGSANDDKQDAPHYIGTLDMTAFGSLFHEDDFELNFGTAK ...	$\#(D, E) = 1$
.....	
VCKITPHAPHKSHPDGVYGTSGSANADRQDAPNYIGTLDMTAFGSLFHEDEFELTFGTTK ...	$\#(N, F) = 0$
KLNELIPTRLDRKGLQSGGKVDRYQDEKYRKGVSPLYFKKSHARKLAGSLTSDAITTLVRA ...	$\#(D, E) = 3$
.....	
RVSDLYGIRLERAGLQSGGKLARYVEASLTTHGLAYNMASTRLLQGAHTGDASDGLVKT ...	$\#(N, F) = 1$
PKNDSHTQVKEGTEQTFVLPKAHAASKLVEDLLGAGVDSKPNGAYTQESDPSSVPEGVTD ...	$\#(D, E) = 5$
:. . . : : . . . : . . . : . . . : . . . : . . . : . . . : . . . : . . . : . . .	
PQFEGFTTGKDGAPLAAVQKQYHATVMFIVMMGGFAVEQKGFGRGSDKDPCHTSHGLLE ...	$\#(N, F) = 2$

Example: Joint Frequencies at Different Divergence Times



Dealing with Different Divergence Times

BLOSUM way: Accept, limit and mix

- Pool confirmed alignments with relatively high identity (e.g., BLOSUM62: at least 62% sequence identity in alignment)
- Mix alignments with identity above threshold to estimate J , P , π , S as above.

Dealing with Different Divergence Times

BLOSUM way: Accept, limit and mix

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PAM way (Dayhoff): Strictly limit, slightly mix, normalize, extrapolate

- Pool confirmed alignments with very high identity (e.g., $> 85\%$)
- Estimate transition probability matrix $P = (P_{ij})$
- Normalize P such that the overall amount of change is 1%:
$$\sum_{i \neq j} J_{ij} = \sum_i \pi_i \sum_{j \neq i} P_{ij} \stackrel{!}{=} 1/100 \text{ (1 PAM = percent accepted mutations)}$$
- Set $P_{ii} := 1 - \sum_{j \neq i} P_{ij}$
- Defines a convenient PAM “unit of evolution”, extrapolate to longer times.

More on Markov Processes

Long-term transition probabilities

- $P = P^{(1)}$: one-step (1 PAM) transition probabilities
- What happens in 120 PAM (time period $120\times$ longer than 1 PAM)?

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- To move $i \rightarrow j$ in 2 PAM, move $i \mapsto k \mapsto j$ for some k :
 $P_{ij}^{(2)} = \sum_k P_{ik}^{(1)} \cdot P_{kj}^{(1)}$ or $P^{(2)} = P^{(1)} \cdot P^{(1)}$.

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 $P_{ij}^{(2)} = \sum_k P_{ik}^{(1)} \cdot P_{kj}^{(1)}$ or $P^{(2)} = P^{(1)} \cdot P^{(1)}$.
- In general, $P^{(s+t)} = P^{(s)} \cdot P^{(t)}$ (Chapman-Kolmogorov)
- Therefore, $P^{(t)} = P^t$, the t -th power of P ($t \in \mathbb{N}$), and $P^{(0)} = P^0 = \text{Id}$.

Dayhoff's Extrapolation Method

PAM Matrix Family

- Remember: $P = P^{(1)}$ was created from mixed closely related alignments, artificially normalized to 1 PAM.
- Now $P^{(120)} = P^{120}$ extrapolates to time span 120 PAM.
- Limits: $P^{(0)} = \text{Id}$ (no change), and $P_{ij}^{(\infty)} = \pi_i \cdot \pi_j$.

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Disadvantages

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- Rare replacements are too infrequent to resolve transition probabilities accurately.
- Errors in 1 PAM matrix are magnified in the extrapolation to 120 PAM.

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- Errors in 1 PAM matrix are magnified in the extrapolation to 120 PAM.

⇒ Design method to utilize alignments of varying divergence (Tobias Müller, ca. 2000)

Continuous-Time Markov Processes

Rate matrices

- For real numbers p , we have $p^n = p \cdot \dots \cdot p$ (n times).
Generalize to real exponents $t \in \mathbb{R}$ by $p^t = \exp(t \cdot \log p)$.
- Can it be done for matrices, too?

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Generalize to real exponents $t \in \mathbb{R}$ by $p^t = \exp(t \cdot \log p)$.
- Can it be done for matrices, too?
- **Definition:** A matrix Q such that $\exp(Q) = P$ is called **rate matrix** or **infinitesimal generator** of the Markov process, or **matrix logarithm**.
- Does not exist for every matrix, but does for positive definite matrices.
Compute by diagonalization and taking logarithm of each (positive) eigenvalue.
- Matrix exponential for square matrices defined by power series

$$\exp(Q) := \text{Id} + Q + Q^2/2 + \dots + Q^k/k! + \dots$$

- Property: $\exp(tQ) \cdot \exp(sQ) = \exp((s + t)Q)$

Understanding the Rate Matrix

Linear approximation for small $t > 0$

- $P^t = \exp(tQ) = \text{Id} + tQ + t^2 Q^2/2 + \dots + t^k Q^k/k! + \dots \approx \text{Id} + tQ$
- Also, $Q = \lim_{t \searrow 0} (P^{(t)} - \text{Id})/t = P'(0)$
- Therefore, Q contains the rates describing how fast P_{ij}^t changes near $t = 0$.

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Some Properties

- Valid rate matrix has $Q_{ij} > 0$ for $i \neq j$ and $Q_{ii} < 0$ for all i .
- Zero row sums: $\sum_j Q_{ij} = 0$ or $Q_{ii} = -\sum_{j \neq i} Q_{ij} < 0$

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Alternative scaling or calibration

- So far: Scale Q such that $P = \exp(Q)$ has 1 PAM (amount of change)
- Alternative: Scale Q such that $\sum_i \pi_i \sum_{j \neq i} Q_{ij} = 1/100$, unit of 1 PEM (percent expected mutation events)

Different Expressions for the Rate Matrix

How to obtain Q from the family $P(t)$?

$$Q = \log(P(t))/t \quad (\text{any } t > 0)$$

$$Q = \alpha \cdot \text{Id} - \left(\int_0^{\infty} e^{-\alpha t} P(t) dt \right)^{-1} \quad (\text{any } \alpha > 0)$$

The integral is called the **resolvent** (or **Laplace transform**) of $P(t)$, $t > 0$.

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Why is the resolvent representation useful?

The resolvent expression integrates (in a weighted manner) over **all** times t .
We can use alignments of different degrees of divergence and adjust weighting via α .

Estimation of Q based on the Resolvent Expression

- 1 Start with an initial rate matrix Q and pairwise alignments (A_i)
- 2 Calculate empirical transition matrix $P_{(i)}$ from A_i for all i
- 3 Estimate divergence time t_i for A_i using existing rates Q
- 4 Combine different $P_{(i)}^{(t_i)}$ with approximately equal t_i
(fewer time points, but better $P^{(t)}$ estimates at each time point t)
- 5 Estimate the resolvent $R_\alpha = \int_0^\infty e^{-\alpha t} P^{(t)} dt \approx \sum_t e^{-\alpha t} P^{(t)}$
for different $\alpha > 0$
- 6 Select “best” parameter α^* by Maximum-Likelihood-like procedure
- 7 Set $Q := \alpha^* \cdot \text{Id} - R_{\alpha^*}^{-1}$
- 8 Iterate steps 3 – 7 until Q converges

Estimating the Divergence Time

Problem

Given alignment A (yielding empirical transition matrix $\hat{P} = \hat{P}^{(t)}$); rate matrix Q ,
estimate divergence time t , such that $\exp(tQ) \approx \hat{P}$

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Approach: Maximum Likelihood

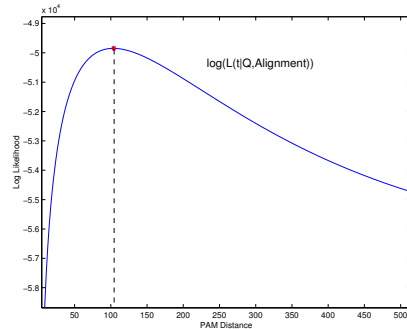
Probability of observing A ,
given Q and time t :

$$\mathbf{P}(A \mid Q, t) = \prod_{i,j} (\exp(tQ)_{ij})^{\hat{P}_{ij}}.$$

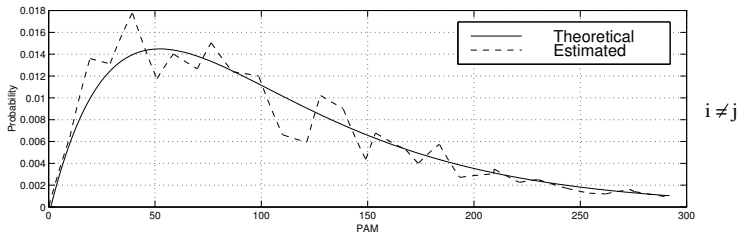
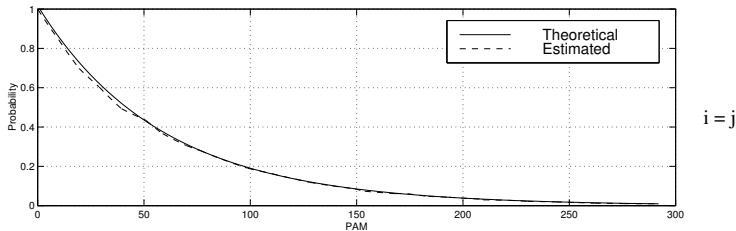
Log-likelihood of t :

$$\log L(t \mid Q, A) = \sum_{i,j} \hat{P}_{ij} \cdot \log[\exp(tQ)_{ij}]$$

→ maximize



Integrating Different $P^{(t)}$ with the Resolvent $\int e^{-\alpha t} P_{ij}^{(t)} dt$



Picking Parameter α and Other Details

Interpretation of α

- Controls speed of exponential decay of weighting: $\sum_t e^{-\alpha t} P(t)$
- Small α : High weight to large divergence times
Large α : Small weight to large divergence times
- Pick α to let $e^{-\alpha t}$ fit amount of alignment data at each time t .
- Can use a maximum-likelihood-like approach or curve fitting

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Starting with an initial Q

- Initially, choose Q such that all rates are equal, calibrate to 1 PEM or or 1 PAM.
- Gives approximate divergence time estimates for first iteration.

Again: Resolvent Estimation of Q

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Families of Score Matrices

Score matrices have a time parameter t

$$\tilde{S}_{ij}^{(t)} = \log_2 \frac{J_{ij}^{(t)}}{\pi_i \cdot \pi_j} = \log_2 \frac{\pi_i P_{ij}^{(t)}}{\pi_i \cdot \pi_j} = \log_2 \frac{\exp(tQ)_{ij}}{\pi_j} \quad [\text{bits}]$$

- PAM family indexed by t (in PAM units), Dayhoff method
- VTML family indexed by t (in PAM units), resolvent method
- BLOSUM family indexed by percent identity (no rate matrix)

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- VTML family indexed by t (in PAM units), resolvent method
- BLOSUM family indexed by percent identity (no rate matrix)

Non-symmetric score matrices?

- So far, score matrices were symmetric, but sequence roles in alignments may differ.
- Search with a transmembrane domain (τ) in a general protein database (π):

may want $\tilde{S}_{ij}^{(t)} = \log_2 \frac{J_{ij}^{(t)}}{\tau_i \cdot \pi_j}$ (e.g., SLIM matrix family, Müller et al., 2001)

Summary

Score Matrices

- Joint frequencies J , transition probabilities P , marginal probabilities π
- (Scaled, rounded) Log-odds scores S
- Evolutionary time units: PAM, PEM
- Markov processes, Chapman-Kolmogorov equation $P^{(s+t)} = P^{(s)} \cdot P^{(t)}$
- Rate matrix Q and time- t transition matrices: $P^{(t)} = P^t = \exp(tQ)$
- Resolvent (Laplace transform) method allows estimation of Q from alignments of varying divergence.
- Symmetric general-purpose vs. (perhaps non-symmetric) special-purpose matrices

Possible Exam Questions

- Define joint frequencies J , transition probabilities P , marginal probabilities π .
- Describe how to compute log-odds scores.
- Explain how the BLOSUM and PAM matrix families were constructed.
- What is a rate matrix?
- Define the evolutionary time units 1 PAM and 1 PEM.
- How can Q be expressed in terms of P or all $P^{(t)}$?
- How can you estimate the divergence time of an observed alignment?
- What are the advantages of the resolvent method for estimating Q ?
- Are score matrices symmetric? Why? When not?